A NOTE ON ESTIMATION IN DOUBLE SAMPLING

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Let \overline{X}_t be the finite population means for variates x_t (t=0, 1, 2). Let \overline{x}'_2 be a first-phase sample (SRSWOR, of size n') mean on x_2 and \overline{x}_t be a second-phase sub-sample (SRSWOR, of size n < n') means on x_t (t=0, 1, 2). For $R=\overline{X}_0/\overline{X}_1$ and for $P=\overline{X}_0\overline{X}_1$, Singh [1] gave double-sample estimators $\hat{R}_1 = \hat{R}\overline{x}_2/\overline{x}_2$, $\hat{R}_2 = \hat{R}\overline{x}_2/\overline{x}_2$ and $\hat{P}_1 = \hat{P}\overline{x}_2/\overline{x}_2$, $\hat{P}_2 = \hat{P}_2\overline{x}_2/\overline{x}_2$ respectively (where $\hat{R}=\overline{x}_0/\overline{x}_1$, $\hat{P}=\overline{x}_0\overline{x}_1$). Their asymptotic mean square errors (*MSE*), to the first order of approximation, dervied using delta method, are vide Singh [1]:

$$MSE(\hat{R}_{t}) = M + fR^{2}c_{2}^{2} [1 - 2(-1)_{c}^{t}], \quad t = 1, 2$$

$$MS(E\hat{P}_{t}) = M_{1} + fP^{2}c_{2}^{2} [1 - 2(-1)^{t}c^{*}], \quad t = 1, 2$$

where $M = MSE(\hat{R}), M_{1} = MSE(\hat{P}), f = \left(\frac{1}{n} - \frac{1}{n'}\right),$

 c_t =coefficient of variation of $x_t(t=0, 1, 2)$,

 ρ_{tt} = correlation coefficient between x_t and $x_{t'}$,

$$(t, t'=0, 1, 2), c=\rho_{02}(c_0/c_2)-\rho_{12}(c_1/c_2)$$

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Writing s_{tt} , for the sum-sample (of size *n*)—covariance between x_t and x_t , (for t, t'=0, 1, 2) we propose better alternative estimators respectively for R and P as

$$\hat{R}^* = \hat{R} + a(\bar{x}_2 - \bar{x}_2),$$

 $c^* = \rho_{02}(c_0/c_2) + \rho_{12}(c_1/c).$

where $a = -\frac{\hat{R}}{s_{23}}(s_{02}/\bar{x}_0 - s_{12}/\bar{x}_1)$

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and

$$\hat{p}^* = \hat{p} + b(\bar{x}_2 - \bar{x}_2),$$

where

$$b = -\frac{P}{s_{22}}(s_{02}/\bar{x}_0 + s_{12}/\bar{x}_1).$$

Their asymptotic (to first order of approximation). MSE's are using delta methods again, respectively,

$$M(\hat{R}^*) = M - fR^2 c_2^2 c_2 < MSE(\hat{R}_t)$$

and $M(\hat{P}^*) = M_1 - fP^2 c_2^2 c_2 < MSE(\hat{P}_t)$ for both $t = 1, 2$

AN ILLUSTRATION

For the example considered by Singh (1965), we have

 $MSE(\hat{R}) = k' (0.5318)$ $MSE(\hat{R}_{1}) = k' (0.5.318) - k^{*}(1.42020)$ $MSE(\hat{R}_{2}) = k' (0.5318) - k^{*}(.2728)$ $MSE(\hat{R}^{*}) = k' (0.5318) - k^{*}(.3123)$

where

$$k' = \frac{(N-n)}{nN} kR_2,$$
$$k^* = \left(\frac{1}{n} - \frac{1}{n'}\right) kR^2,$$

 $k = \frac{N}{(N-1)}$ and N is the number of blocks in the population.

From above, we see that \hat{R}^* is the most efficient.

REFERENCE

[1] Singh, M.P. (1965)

: On the estimation of ratio and product of the population parameters. Sankhya, 27 (B), 32-128.